

# Sufficient Statistics for Economic Mobility: When Do Measures Agree?

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# A “proliferation of measures”

- Researchers measure economic mobility

  - ... in many outcomes (e.g., incomes, wealth, education, health, ...)

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- Different measures often produce consistent rankings of mobility (e.g., Katz and Krueger 2017, Berman 2022, Deutscher and Mazumder 2023)
- But not always: long-run US trends (Jácome, Kuziemko & Naidu, 2025); cross-CZ w/in US (Chetty et al., 2014); cross-country (Blanden 2013) 

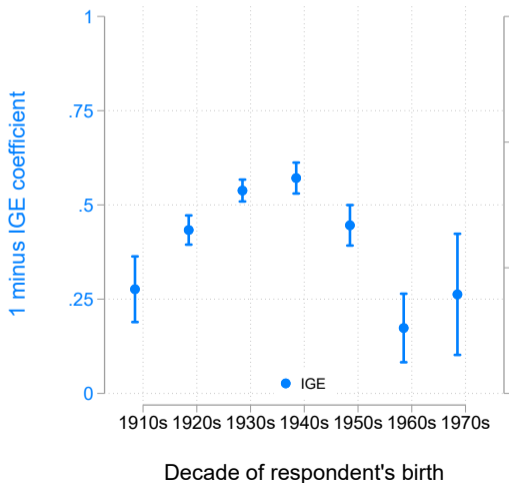
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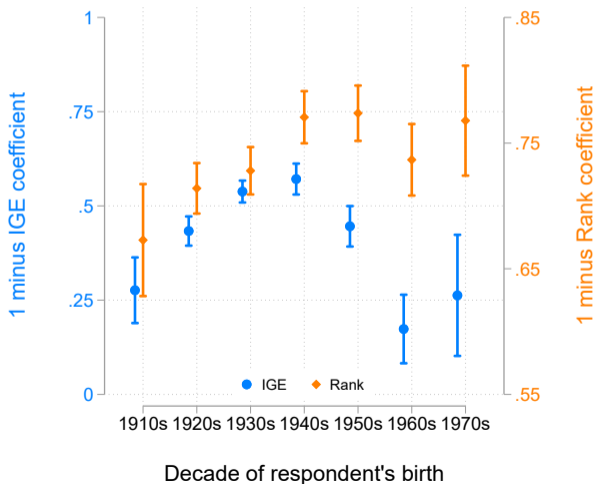
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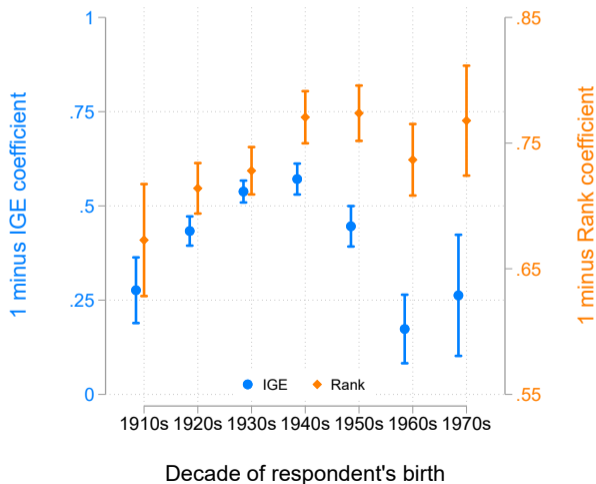
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Two fundamental questions:

1. When *should* we expect measures to agree?
2. What does disagreement *tell us*?



## This Paper Provides Answers

1. Agreement of measures depends on single property of parent-child joint distribution
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  - Fixing marginal distributions of income  $\rightarrow$  relative mobility
3. Changes in any mobility measure = rank-dependence + changing inequality
  - Rank-based measures: only rank-dependence
  - Level-based (incl. IGE): both  $\rightarrow$  potential for disagreement
  - Disagreement reflects:
    - i. different weightings over regions in joint distribution
    - ii. different sensitivity to changing inequality

# Contributions

1. Unifying framework w/ **Concordance** as sufficient statistic for mobility measurement
  - Adds to measure comparisons (Deutscher, Mazumder; Berman), axiomatic approaches (Shorrocks; Fields, Ok; D'Agostino, Dardanoni; Cowell, Flachaire), dual of 2D-inequality problem (Atkinson, Bourguignon; Dardanoni)
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3. Microfounded w/ structural interpretation in Becker-Tomes-Loury models
  - Sorting measures NAM vs PAM in matching (Anderson, Smith; Gola; Boerma, Tsyvinski, Wang, Zhang) and mating (Chiappori, Costa Dias, Meghir, Zhang)

Setup

## Separating Dependence from Marginals: Copulas

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**Sklar's (1959) Theorem:** Any joint distribution decomposes as:

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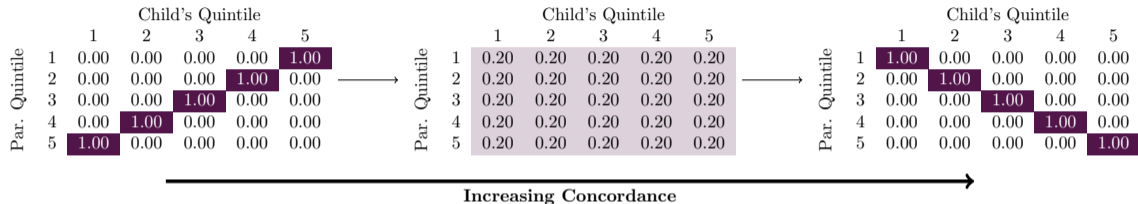
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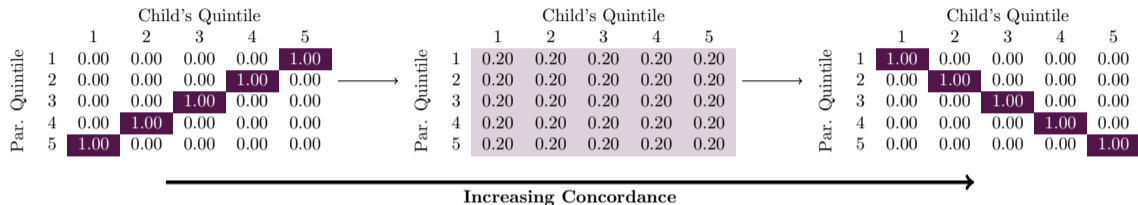
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- Copula encodes *all* rank-dependence between generations  $\perp$  of marginals ▶ Mass points
- Relative mobility measures alternative summaries of copula

# Concordance Order: Visual Intuition ▶ Example II

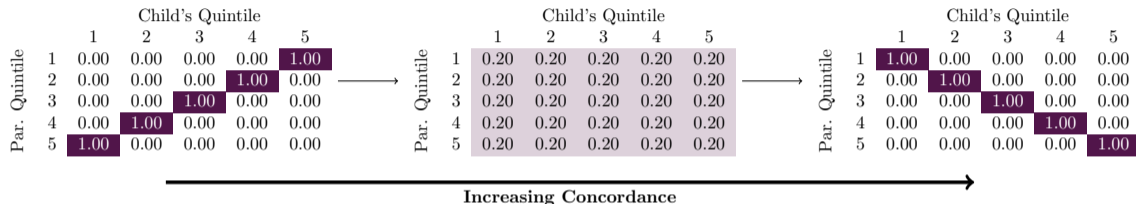


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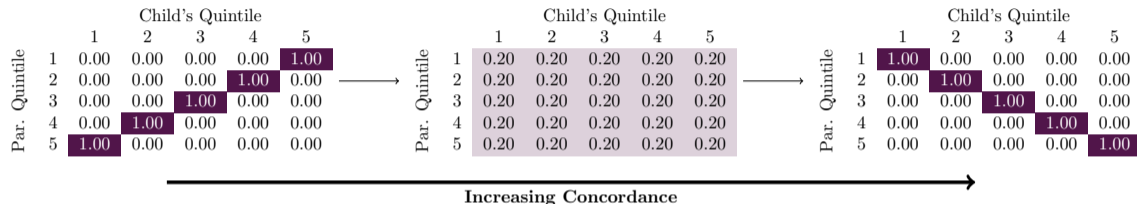
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- Pinned down by dependence parameters in parametric families (e.g.,  $\rho$  in normal)

## The Concordance Order

**Definition:** (Tchen 1980) Economy  $A$  is **more concordant** than economy  $B$  ( $C^A \succeq C^B$ ) if:

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  - Incomplete order over distributions, but covers empirically relevant cases [▶ jump e.g.](#)
    - “intuitive requirement that children from higher-income families are less likely to have lower incomes” (Chetty et al., 2017); “empirically unlikely transition matrices” (Shorrocks, 1978); “all-or-nothing” (Berman, 2022)
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- Higher concordance  $\longleftrightarrow$  **lower mobility**

# Main Results: Sufficiency

# Concordance Orders All Standard Mobility Measures

**Props 1-4:** Fix marginals. If  $C^A \succeq C^B$ , then economy A is *less mobile* than B by:

1. Regression/Correlation-based measures:
  - IGE:  $\beta^A \geq \beta^B$ , Rank-rank slope:  $\rho^A \geq \rho^B$ , Intergenerational correlation
2. “Local” measures:
  - Conditional expected rank  $\rightarrow$  rotates around  $r^* = 0.5$
3. Transition probabilities:
  - diagonal  $\uparrow$ , off-diagonal  $\downarrow$ , Directional mobility
4. Axiomatic measures:
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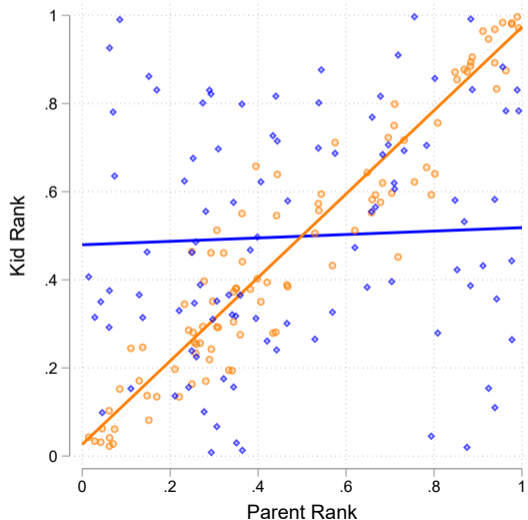
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$\rightarrow$  Measures are **ordinal substitutes**: observe one ranking, know all others

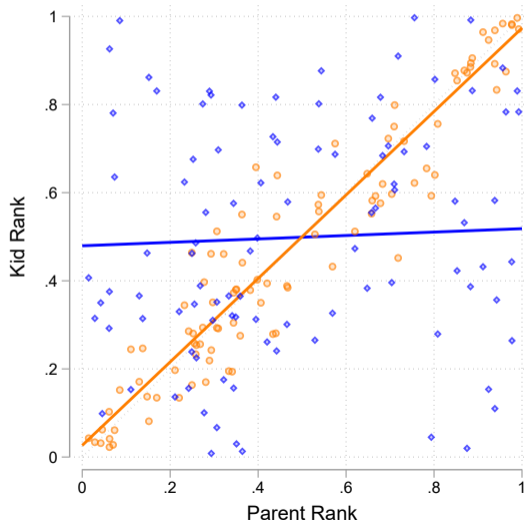
# Why Does Concordance Order These Measures?

## Rank-rank slope



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Mobility Measure:

$$\rho = 12 \iint C(u, v) du dv - 3$$

$C^A \geq C^B$  everywhere

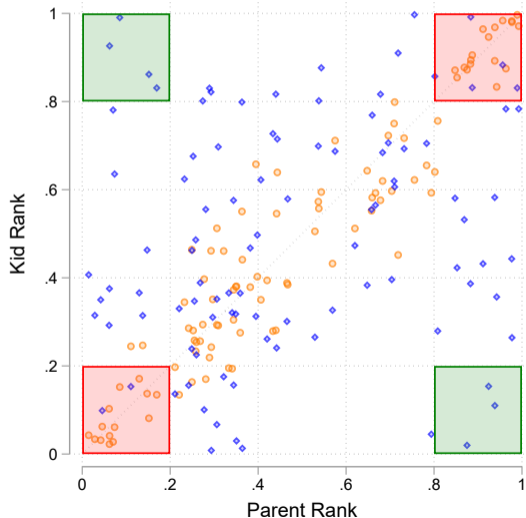
$$\rightarrow \iint C^A \geq \iint C^B$$

$$\rightarrow \rho^A \geq \rho^B$$

(Lehmann 1966; Tchen 1980)

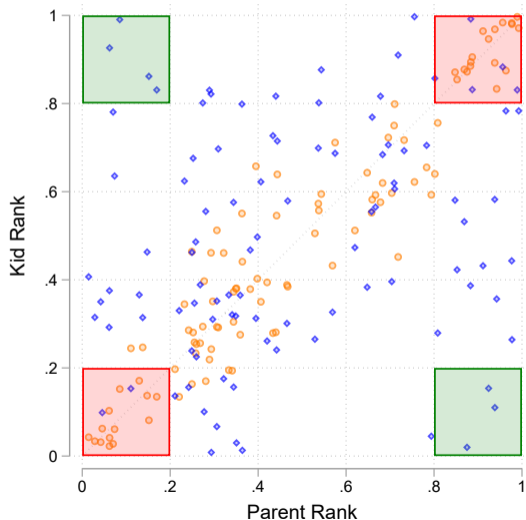
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Mobility Measure:

$$\Pr(R^K > \tau \mid R^P > \tau) = \frac{1 - 2\tau + C(\tau, \tau)}{1 - \tau}$$

If  $C^A \geq C^B$  everywhere

$$\longrightarrow C^A(\tau, \tau) \geq C^B(\tau, \tau)$$

$$\longrightarrow TP^A \geq TP^B$$

(Copula definition, simplest cases!)



# Concordance is Pigou-Dalton for Mobility

- Inequality measurement (Atkinson, 1970)
  1. Elementary operation: Pigou-Dalton transfer
  2. Partial order: Lorenz dominance (SOSD)
  3. Sufficiency: SOSD  $\longrightarrow$  all transfer-sensitive measures agree

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- Mobility measurement (Today)
  1. Elementary operation: **Concordance Exchanges**
    - swap  $Y_i^K$  and  $Y_j^K$  if  $Y_i^K < Y_j^K$ , but  $Y_i^P > Y_j^P \rightarrow$  increase concordance
  2. Partial order: **Concordance**
    - Discuss implications & completions
  3. Sufficiency: Concordance  $\rightarrow$  all exchange mobility measures agree

Analogues in SWF for 2D ineq. (Atkinson-Bourguignon) & Perm. Income ineq. (Dardanoni)

# U.S. Mobility Trends & Formal Test


# Empirical Application: 20<sup>th</sup> Century U.S. Mobility

Datasets: Replicate *Jácome, Kuziemko & Naidu, 2025*

- Pool 15 U.S. surveys (ANES, GSS, PSID, OCG, NLS, ...) with respondent income + father's occupation (large sample of intergenerational links) 
- Impute parental income from Census using occupation  $\times$  race  $\times$  region (using census)
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**Theory predicts:**

- Rank-based measures should **co-move**, but other measures may diverge (rising inequality)

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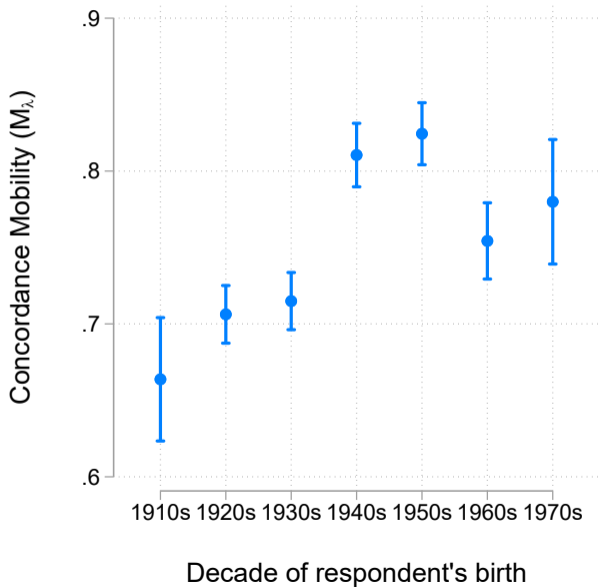
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Return to formal testing procedure

# Concordance Mobility over the 20<sup>th</sup> Century



## Wave-pattern

- Long run Mobility rises over the 20<sup>th</sup> century
- But plateaus/declines after 1960s

# Increasing Mobility? Moment Inequality Test of Concordance Order

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1920	1930	.001	.004	Copulas cross
1930	1940	.052	.001	$C_{1930} \succeq C_{1940}$
1940	1950	.142	.358	Fail to reject
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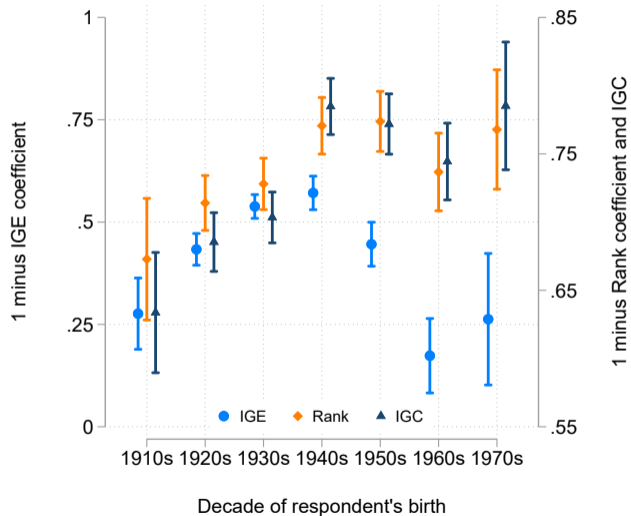
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- Some measure disagreement (paper has many measures) noise, others substantive  
→ Transition Probabilities disagree when copulas cross

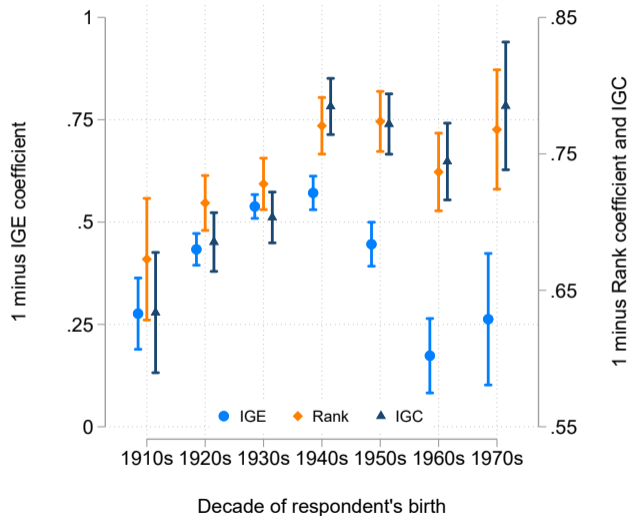
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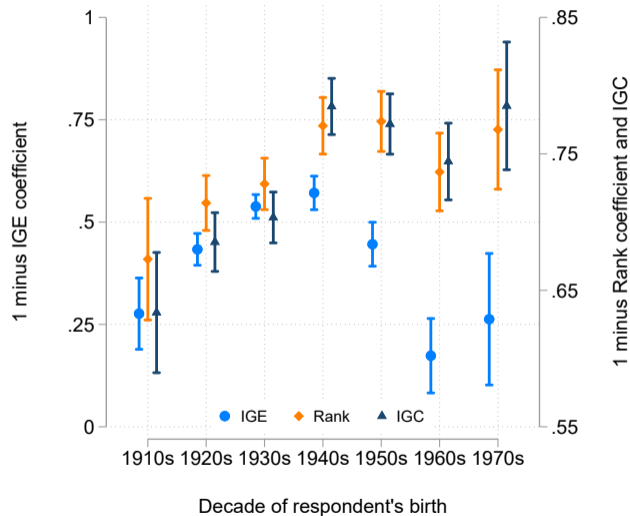
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- IGC tracks rank-rank slope (not generic)
- IGE diverges starting 1950s

Why? IGE adds changing inequality to theory results

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- Level-based measures (IGE): Both components non-zero & potentially opposites
- Order matters  $\longrightarrow$  average over fixing copulas at  $A$  and  $B$   
(Shapley-Owen-Shorrocks decomposition)

# Decomposition: Why Does IGE Diverge?

▶ Other Measures



- **Dependence component:** rising mobility + plateau
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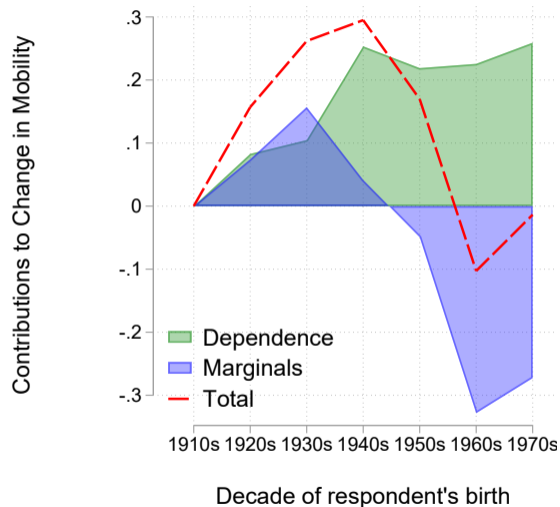


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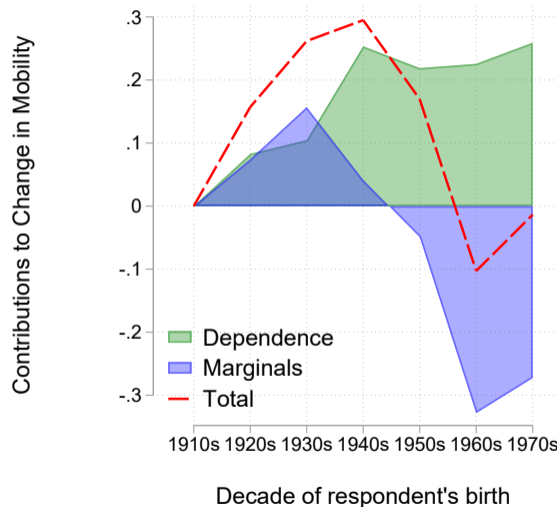
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- Attenuates post-1940 & dominates by 1970
- Disagreement reflects sensitivity to marginals!

# What To Do When Concordance Fails?

## Crossing Copulas: 1920s–30s and 1950s–60s ◀

Concordance is a partial order → not all economies are ranked, e.g.

		Child's Quintile				
		1	2	3	4	5
Par. Quintile	1	0.32	0.24	0.2	0.15	0.08
	2	0.25	0.23	0.21	0.18	0.12
	3	0.19	0.22	0.22	0.21	0.16
	4	0.15	0.18	0.21	0.24	0.22
	5	0.09	0.12	0.16	0.22	0.41

No  
Concordance  
Order  
←————→

		Child's Quintile				
		1	2	3	4	5
Par. Quintile	1	0.58	0.00	0.00	0.14	0.28
	2	0.00	0.86	0.00	0.00	0.14
	3	0.00	0.00	1.00	0.00	0.00
	4	0.14	0.00	0.00	0.86	0.00
	5	0.28	0.14	0.00	0.00	0.58

- Identical rank-correlation (0.32)



## Source of Disagreement

- Measures expressed as

$$M = \int \int \phi_M(Y^K, Y^P) f(Y^K, Y^P) dY^K dY^P,$$

where  $\phi_M(Y^K, Y^P)$ : mobility-measure specific intensity


- TP: Indicator; IGE:  $1/\sigma_P^2(\ln Y^K - \ln \bar{Y}^K)(\ln Y^P - \ln \bar{Y}^P)$

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- Concordance guarantees  $C^A - C^B \geq 0 \rightarrow \Delta M \leq 0 \perp$  of weights 
- Failures  $\rightarrow C^A - C^B$  changes sign on diff. regions, but measures diff. sensitivity
  - Tie-break: **Normative** (intersecting Lorenz curves) or **robustness** choice (sensitivity to  $\epsilon$ -violations)

# What To Report When Concordance Fails

**Prop 6:** Concordance isn't knife-edge!

If  $C^A - C^B$  within  $\epsilon$  of satisfying concordance a measure  $M$  agrees up to  $L_M \cdot \epsilon$

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**Reporting protocol:** use test output to localise the violation

- Report **local rank-rank slopes** on the regions where  $C^A - C^B$  changes sign
- Supplement with transition probabilities at standard thresholds

Lets readers with different preferred measures reach their own conclusions

# Other Settings & Structural Interpretation

No time? Jump to Conclusions

## Other Settings ▶ Skip

1. More measures (follow JEL taxonomy, Deutscher & Mazumder) include absolute measures
  - Lorenz mobility (McGee & Ocampo, 2025), Bartholomew jump measure, Goodman, average non-linear income persistence (Arellano, Blundell, Bonhomme, 2017), Katz & Krueger
2. Concordance order maintained under fixed rank comparisons
  - Ranking children by position in parent distributions
  - Subgroup or subnational regions against population distribution
3. Allow for multi-dimensional mobility
  - Whole lifecycle paths (Audoly, McGee, Ocampo & Paz-Pardo, 2025) or many generations

## Microfoundation: Becker-Tomes-Loury

- Standard dynastic investment model

$$\max_{C^P, I} U(C^P) + \delta \mathbb{E}_\theta [U(C^K)]$$

subject to

$$C^P = Y^P - I, \quad C^K = Y^K = W \left( \underbrace{\theta f(I, H^P)}_{H^K} \right)$$

- Parents choose investment ( $I$ ) and consumption ( $C^P$ ); child consumes ( $C^K$ )
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- $\theta$ : luck shock realized after investment
- When do changes in technology or preferences lead to more mobility?

## Investment Expansions $\longrightarrow$ Larger in Concordance ▶ Example

**Lemma 1:** If investment expands uniformly ▶

$$I^{*A}(Y^P) \geq I^{*B}(Y^P) \quad \forall Y^P \longrightarrow C^A \succeq C^B$$

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1. Higher marginal product of investment,  $f_I^A(I, H^P) \geq f_I^B(I, H^P) \quad \forall (H^P, I)$  or subsidy;
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What about **heterogeneous** policy impacts?

- E.g. school finance equalisation: poor districts  $\uparrow$ , rich districts  $\downarrow$  (Biasi, 2023)
- $\Delta I^*(Y^P)$  crosses zero  $\longrightarrow \Delta C$  changes sign  $\longrightarrow$  **crossing copulas**
- Rationalises 1920s–30s, 1950s–60s?

# Conclusions

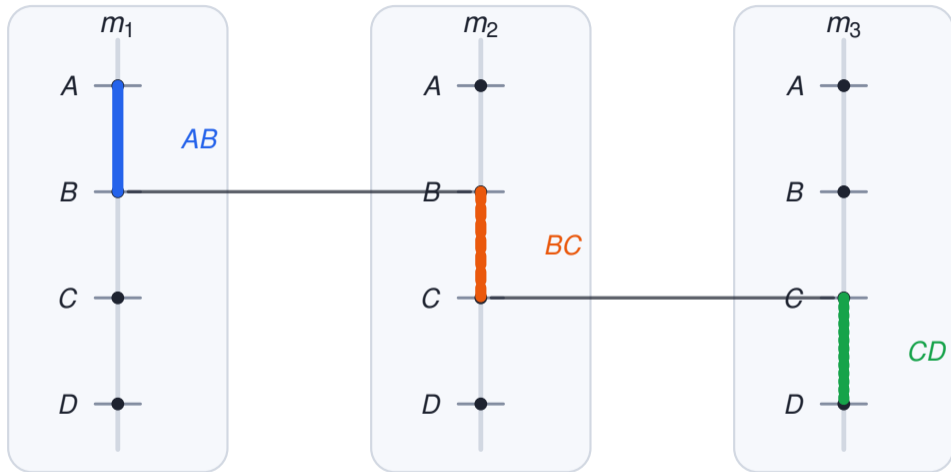
# Implications For Applied Researchers

A single property—**concordance**—determines when standard mobility measures agree

## For applied researchers:

1. **When concordance holds:** measures are substitutes
  - Pick by data quality, not axiom: transitions for noisy data, IGE for clean panels, rank-rank for cross-country
2. **When it fails:** disagreement is informative, not noise
  - Test it; report local rank-rank slopes; Lipschitz bounds for  $\epsilon$ -violations
3. **Decomposition as diagnostic:** rank-dependence vs. changing inequality  
(reframes IGE-vs-rank-rank divergence)

# Chaining Measures



Gives full comparison of A, B, C and D in *any* measure

# Details

e.g. from bunching, top-coding, censoring, genuine discreteness  $\rightarrow$  no restrictions on  $F_K$  or  $F_P$ , but Copula 'indeterminate'

Solutions:

### 1. Parametric copulas

- Does violence to discrete data
- Restrictive

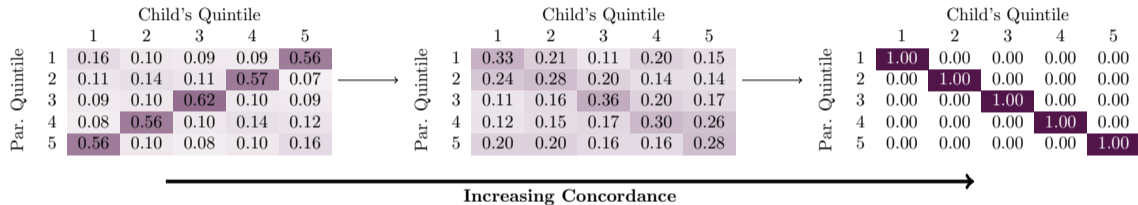
### 2. Carley (2002) bounds on copulas

- A lot more math, but flexible
- Worst-case bounds may not be tight

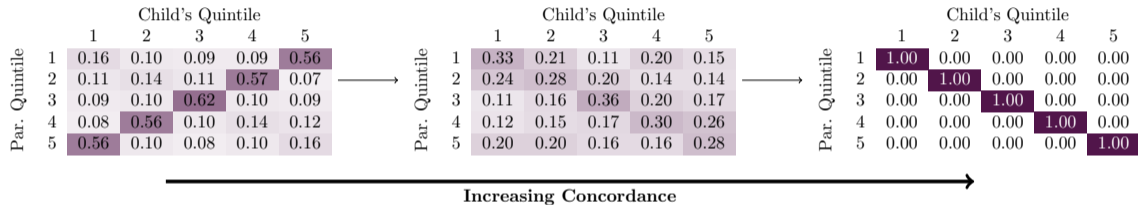
### 3. Pick an extension: checkerboard or Piecewise bilinear

- Propositions 11, 13 in Genest and Nešlehová (2007) show the unique extension preserving concordance order of the original joint-distributions

# Concordance: Visual Intuition [← Back](#)



# Concordance: Visual Intuition [← Back](#)



- Not just extreme cases

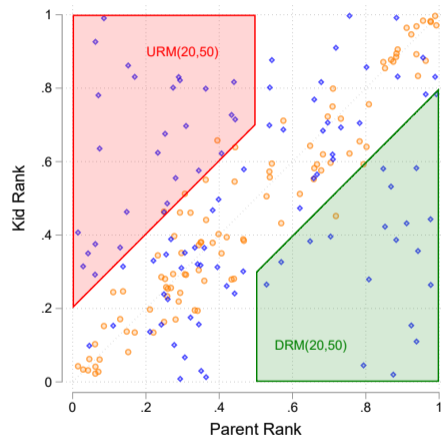
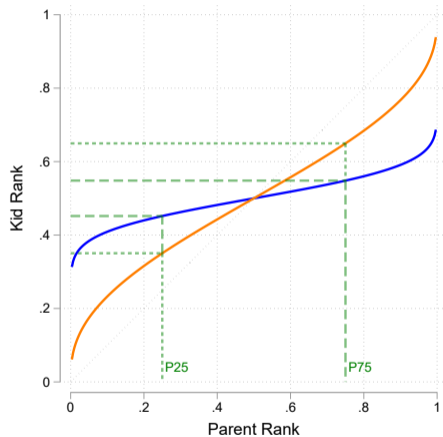
Some results require extra structure

1. Single crossing of conditional densities and continuity (sufficient: TP2 or MLRP)
  - For Non-parametric CER rotation
2. Smooth Anti-Diagonal to Diagonal Exchanges:

$$\Delta C(u_1, v_1) + \Delta C(u_2, v_2) \geq \Delta C(u_2, v_1) + \Delta C(u_1, v_2).$$

- Implies Concordance Order
- Directional Mobility & Shorrocks

# Why Does Concordance Order These Measures? Illustration III [← Back](#)



## Extension: Proxies & Measurement Error [◀ Back](#)

**Prop 5:** Observe proxy  $E^g = H_g(R^g, \varepsilon^g)$  for generation  $g$ ,  $H_g$  increasing, and  $\varepsilon$  independent of ranks. Let  $\tilde{C}$  denote copula for  $(E^K, E^P)$ . **Concordance ordering preserved:**

$$C^A \succeq C^B \quad \longrightarrow \quad \tilde{C}^A \succeq \tilde{C}^B$$

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2. Income and Education as skill proxies:  $S \longrightarrow E, Y$

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- Covers flexible measurement error cases  $\tilde{Y} = E = H_g(R^g, \varepsilon^g)$  (Corollary 1) ▶

1. (non-)Classical measurement error in (log-)incomes (Bound, Brown, Duncan & Rogers 1994)

2. Measurement error in ranks (Chetverikov & Wilhelm 2023)

3. Life cycle biases (Kitagawa, Nybom, & Stuhler 2018)

## Proxies & Meas Error: Functional Forms [← Back](#)

Proxies:

1. Human capital determines incomes  $Y = H(E) + \varepsilon$  with  $E = H^{-1}(Y - \varepsilon)$
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Measurement error:

1. Classical measurement error in incomes

$$\tilde{Y}^K = F_K^{-1}(R^K) + \varepsilon^K, \quad \tilde{Y}^P = F_P^{-1}(R^P) + \varepsilon^P$$

2. Measurement error in ranks

$$\tilde{R}^K = F_{\tilde{K}} \left( F_K^{-1}(R^K) + \varepsilon^K \right), \quad \tilde{R}^P = F_{\tilde{P}} \left( F_P^{-1}(R^P) + \varepsilon^P \right)$$

3. Life Cycle Biases

$$\tilde{Y}^K = G_K \left( F_K^{-1}(R^K) + \mu_{\varepsilon^K} + \sigma_{\varepsilon^K} \varepsilon^K \right), \quad \tilde{Y}^P = G_P \left( F_P^{-1}(R^P) + \mu_{\varepsilon^P} + \sigma_{\varepsilon^P} \varepsilon^P \right)$$

## Measure specific weights [◀ Back](#)

- Measures expressed as

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where  $\phi_M(Y^K, Y^P)$ : mobility-measure specific intensity

- For fixed marginals, any twice-differentiable measure:  $M_M^A - M_M^B$

$$= \int_0^1 \int_0^1 W_M(R^K, R^P; F^K, F^P) [C^A(R^K, R^P) - C^B(R^K, R^P)] dR^K dR^P$$

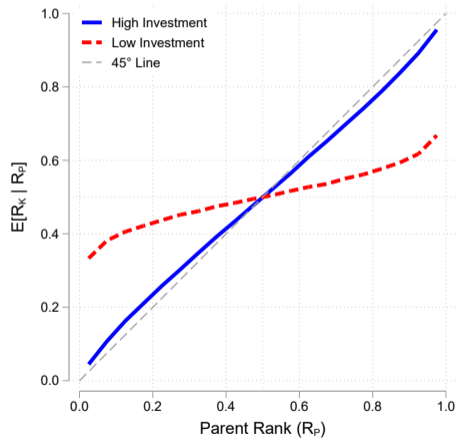
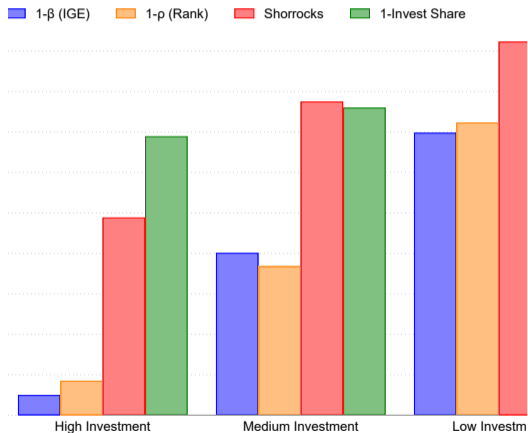
where

$$W_M(R^K, R^P) = \frac{\phi_{M,12}(F_K^{-1}(R^K), F_P^{-1}(R^P))}{f_K(F_K^{-1}(R^K)) f_P(F_P^{-1}(R^P))}$$

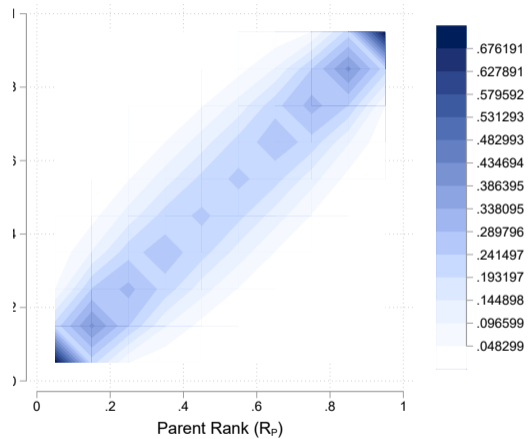
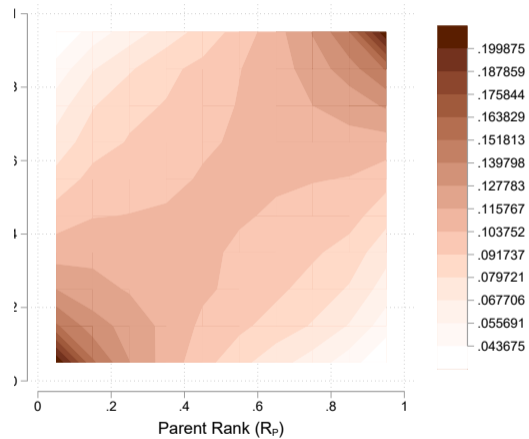
- Works for non-continuous cases also

# Investment Expansions Example [◀ Back](#)

## Log-utility and Cobb-Douglas Production



# Investment Expansions Example [◀ Back](#)



## Proof Sketch: Lemma (Investment $\longrightarrow$ Concordance) [◀ Back](#)

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$$H^{K,A} \mid Y^P = y \succeq_{\text{FOSD}} H^{K,B} \mid Y^P = y \quad \forall y$$

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- But CDF not common, 3 cases:

1. *Rank-dependent wages*: Integrate over  $Y^P \Rightarrow (R^K, Y^P)^A \succeq (R^K, Y^P)^B$
2. *Common wage function*: FOSD of child incomes + decomposition theorem  $\longrightarrow$  concordance  $\uparrow$
3. *Increased returns*: Analogous reasoning applies

## Proof Sketch: Corollary (Policy $\longrightarrow$ Investment) [◀ Back](#)

Define value function for investment  $I$  and 'policy'  $\lambda$ :

$$V(I; \lambda) = U(C^P) + \delta_\lambda \mathbb{E}_\theta [U(W_\lambda(H^K))]$$

subsume subsidy or transfer into  $f$

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**Key condition (Topkis):** Increasing differences holds if

$$\frac{\partial^2 V}{\partial I \partial \lambda} = \frac{\partial}{\partial \lambda} \left[ \delta_\lambda \frac{\partial f_\lambda}{\partial I} \mathbb{E}_\theta [U'(W_\lambda) W'_\lambda] \right] \geq 0$$

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$$\frac{\partial^2 V}{\partial I \partial \lambda} = \frac{\partial}{\partial \lambda} \left[ \delta_\lambda \frac{\partial f_\lambda}{\partial I} \mathbb{E}_\theta [U'(W_\lambda) W'_\lambda] \right] \geq 0$$

**Verification:**

- Production function  $f$  is supermodular
- Wage schedule  $W_\lambda$  is strictly increasing
- Standard regularity on GE effects (e.g., multiplicative shifts)

## Proof Sketch: Corollary (Policy $\longrightarrow$ Investment) ◀ Back

Define value function for investment  $I$  and 'policy'  $\lambda$ :

$$V(I; \lambda) = U(C^P) + \delta_\lambda \mathbb{E}_\theta [U(W_\lambda(H^K))]$$

subsume subsidy or transfer into  $f$

**Key condition (Topkis):** Increasing differences holds if

$$\frac{\partial^2 V}{\partial I \partial \lambda} = \frac{\partial}{\partial \lambda} \left[ \delta_\lambda \frac{\partial f_\lambda}{\partial I} \mathbb{E}_\theta [U'(W_\lambda) W'_\lambda] \right] \geq 0$$

**Verification:**

- Production function  $f$  is supermodular
- Wage schedule  $W_\lambda$  is strictly increasing
- Standard regularity on GE effects (e.g., multiplicative shifts)

**Conclusion:**  $\lambda' < \lambda \implies I^*(Y^P; \lambda') \leq I^*(Y^P; \lambda)$

# Jácome, Kuziemko & Naidu (2025) Replication [← Back](#)

## 15 Pooled Surveys:

- General Social Survey (GSS), American National Election Survey (ANES)
- Panel Study of Income Dynamics (PSID), Occupational Changes in a Generation (OCG)
- National Longitudinal Surveys (NLS): Mature Women, Older Men, Younger Women
- National Survey of Black Americans, National Fertility Survey, others

## Sample Construction:

- Inclusion: US-born, ages 30–50, non-missing family income + father's occupation
- Race: White and Black only (other races < 1% pre-1950)
- Weights: Reweighted to census race  $\times$  sex shares per birth decade

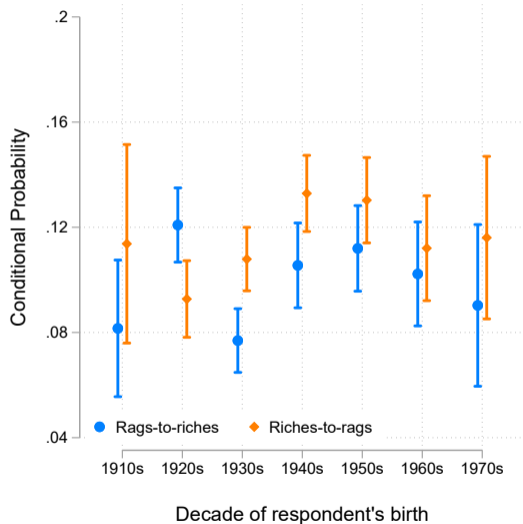
## Parental Income Imputation:

- Predict  $\hat{Y}^P$  using 28 harmonized occupation categories  $\times$  race  $\times$  South
- Sources: 1901 Cost of Living Survey, 1936 Expenditure Survey, 1940–90 Censuses
- Matched to respondent's  $\approx$ 10th birthday; special adjustments for farmers

## Sample Sizes by Birth Decade:

1910s	1920s	1930s	1940s	1950s	1960s	1970s
4,641	11,857	11,426	10,288	9,531	5,691	2,668

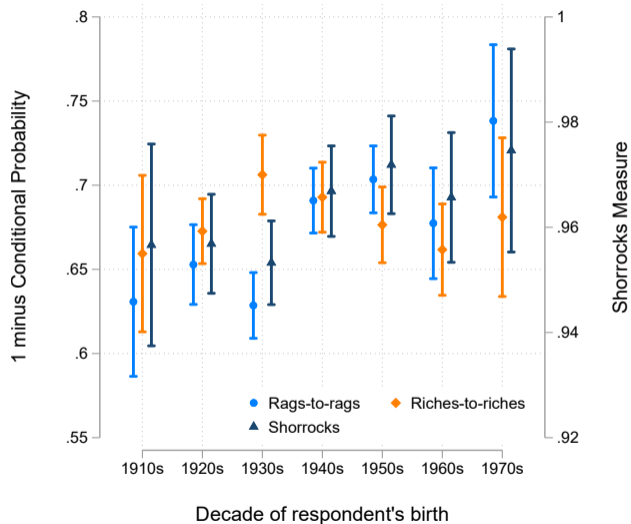
# Transition Matrix Measures Co-Move [◀ Back](#)



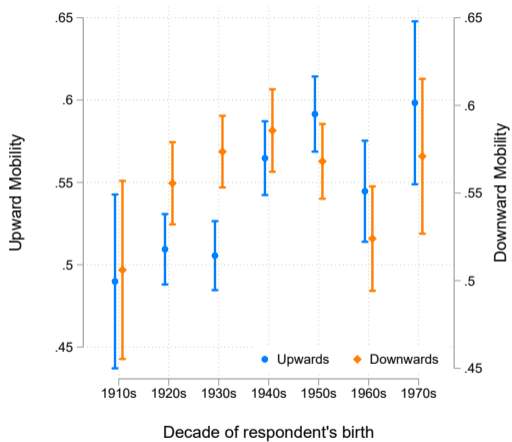
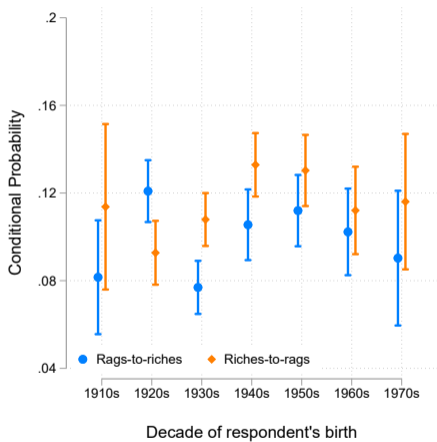
- Quintile Transition Matrix

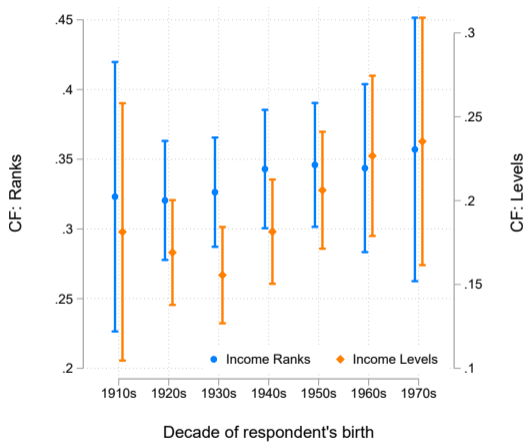
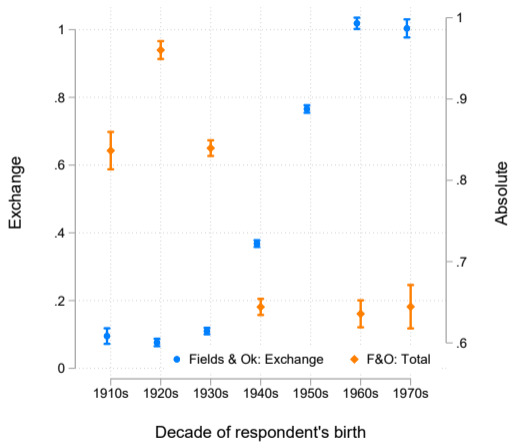
- Q1→Q5 & Q5→Q1

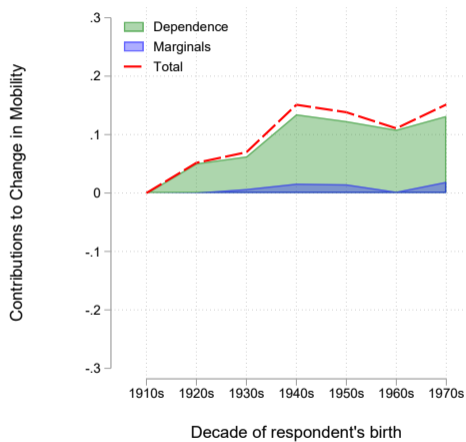
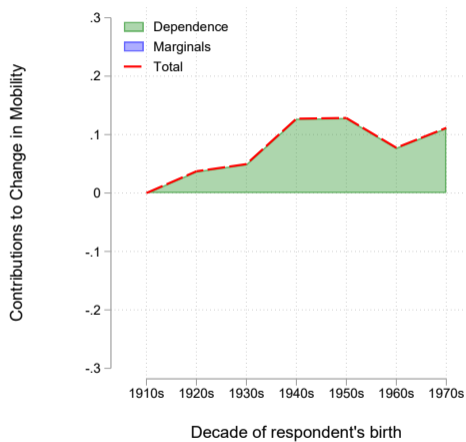
# Transition Matrix Measures Co-Move [◀ Back](#)

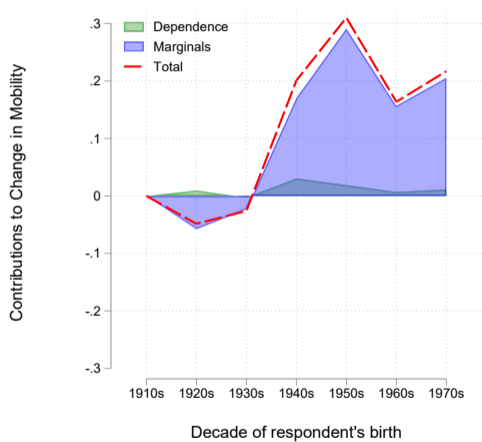
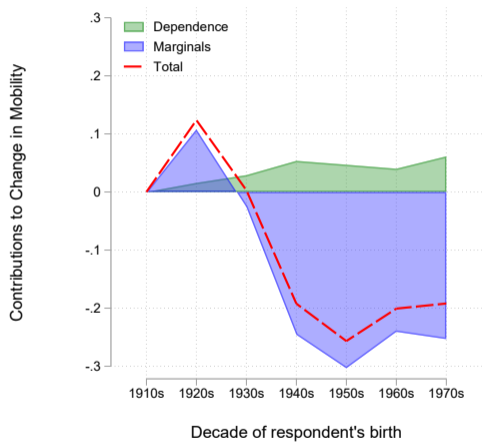


- Quintile Transition Matrix
- Q5→Q5 & Q1→Q1 & Trace









## Testing The Concordance Order [◀ Back](#)

Set up as a moment inequality testing procedure (Andrews and Soares, 2010)

- Moments on a grid  $\mathcal{G}$ :

$$m_g = C_A(u_g, v_g) - C_B(u_g, v_g), \quad g = 1, \dots, G.$$

- Null hypothesis of dominance

$$H_0: m_g \geq 0 \text{ for all } g = 1, \dots, G \quad \text{vs.} \quad H_1: m_g < 0 \text{ for some } g.$$

- Test statistic is the minimum of the studentized moments:

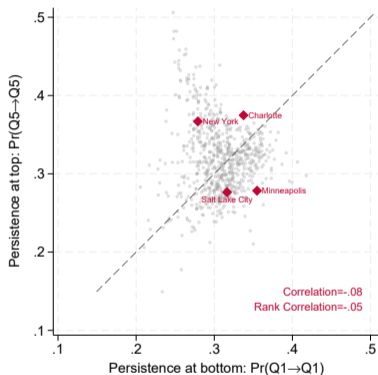
$$T_n = \min_{g=1, \dots, G} \frac{\hat{m}_g}{\hat{s}_g},$$

w/ bootstrap standard error  $\hat{s}_g$

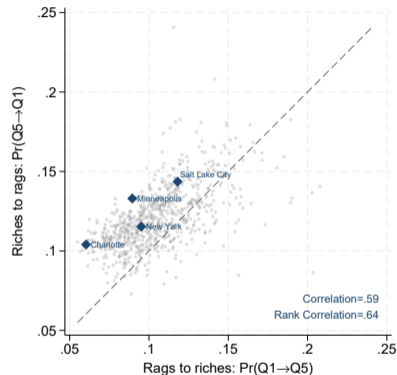
- Test  $A \succeq B$  and  $B \succeq A$

# Spatial Mobility: Measure Disagreement Across CZs [◀ Back](#)

Chetty et al. (2014) CZ transition matrices, remapped to within-CZ copulas



Q1→Q1 vs. Q5→Q5 persistence

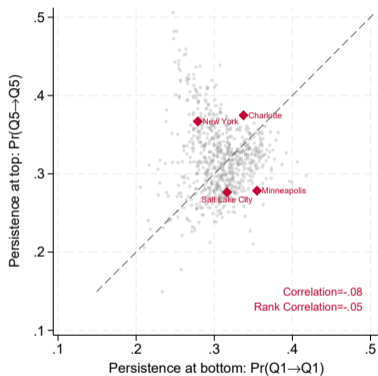


Q1→Q5 vs. Q5→Q1 tail mobility

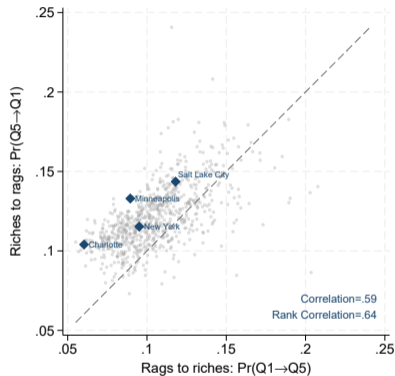
- Salt Lake > Charlotte on every measure...
- ...but Salt Lake vs. NYC, NYC vs. Minneapolis: depends on measure

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Q1→Q5 vs. Q5→Q1 tail mobility

- Salt Lake  $\succ$  Charlotte on every measure...
  - ...but Salt Lake vs. NYC, NYC vs. Minneapolis: depends on measure
- Persistence top-quartile  $\neq$  in **91%** of CZs; **61%** for the tail-mobility

# A Simple Illustrative Example [◀ Back](#)

Separating Dependence & Inequality

Assume Incomes jointly Log-Normal (closed form)

$$\begin{pmatrix} \ln Y^K \\ \ln Y^P \end{pmatrix} \sim N \left[ \begin{pmatrix} \mu_K \\ \mu_P \end{pmatrix}, \begin{pmatrix} \sigma_K^2 & \rho\sigma_K\sigma_P \\ \rho\sigma_K\sigma_P & \sigma_P^2 \end{pmatrix} \right]$$

# A Simple Illustrative Example [◀ Back](#)

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$$\begin{aligned} \text{- Changes: } \Delta\beta = & \underbrace{(\rho^B - \rho^A) \cdot \frac{\sigma_K^B}{\sigma_P^B}}_{\text{change in rank correlation}} + \underbrace{\rho^A \cdot \left( \frac{\sigma_K^B}{\sigma_P^B} - \frac{\sigma_K^A}{\sigma_P^A} \right)}_{\text{change in inequality}} \end{aligned}$$

# A Simple Illustrative Example ◀ Back

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- Rising inequality ( $\sigma_K/\sigma_P \uparrow$ ) can *dominate* IGE even as rank dependence *weakens* ( $\rho \downarrow$ )